

MATH 30440 - SPRING 2010 - HOMEWORK 8 SOLS

7.1) Suppose readings are x_1, \dots, x_n .

Joint density at these readings is

$$f(\theta) = \begin{cases} e^{-(x_1-\theta)} e^{-(x_2-\theta)} \dots e^{-(x_n-\theta)} & \text{if all } x_i \geq \theta \\ 0 & \text{otherwise} \end{cases}$$

In interesting (non-zero) region,

$$f(\theta) = e^{-\sum x_i + n\theta}$$

$$\log f(\theta) = -\sum x_i + n\theta$$

Derivative = n ; so $\log f(\theta)$ always increasing.

We make it as large as possible by choosing θ as large as possible.

Since we have to have $x_i \geq \theta \forall i$,

Choose $\theta =$ smallest number among x_1, \dots, x_n .

7.3) Joint density of n independent copies of Normal (μ, σ^2) at readings (x_1, \dots, x_n) is

$$\left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n e^{-\sum_{i=1}^n \frac{(x_i - \mu)^2}{2\sigma^2}}$$

Log of this is $-n \log \sqrt{2\pi} - n \log \sigma - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$

Derivative: $\frac{-n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2$

$= 0$ when $\sigma^2 = \frac{\sum_{i=1}^n (x_i - \mu)^2}{n}$ ← MLE

$$\begin{aligned} E(\text{MLE}) &= E\left(\frac{\sum_{i=1}^n (x_i - \mu)^2}{n}\right) = \frac{1}{n} \sum_{i=1}^n E((x_i - \mu)^2) \\ &= \frac{1}{n} \sum_{i=1}^n \text{Var}(x_i) \quad // \leftarrow \text{definition!} \\ &= \frac{1}{n} n \sigma^2 \\ &= \sigma^2 \end{aligned}$$

7.7) Best thing we can do, I think, is use

$$\bar{X} = .254$$

$$S = .005$$

as estimates for μ and σ , and calculate

$$P(\text{Normal}(.254, (.005)^2) > .260) \\ + P(\text{ " " " " } < .240)$$

$$\text{ie } P(Z > 1.2) + P(Z < -2.8).$$

No really good way to work in fact
that there are 20 samples.

$$7.8) \bar{X} = 3.1502$$

$$95\% \text{ con int: } \bar{X} \pm 1.96 \frac{.1}{\sqrt{5}}$$

$$99\% \text{ con int } \bar{X} \pm 2.58 \frac{.1}{\sqrt{5}}$$

$$7.9) \quad \bar{X} = 11.48$$

$$a) \text{ 95\% con int: } \quad \bar{X} \pm 1.96 \frac{.08}{\sqrt{10}}$$

$$b) \text{ 95\% lower: } \quad \left(-\infty, \bar{X} + 1.645 \frac{.08}{\sqrt{10}} \right)$$

$$c) \text{ 95\% upper: } \quad \left(\bar{X} - 1.645 \frac{.08}{\sqrt{10}}, \infty \right)$$

$$7.11) \quad a) \quad X_{n+1} = \text{Normal}(\mu, 1)$$

$$\bar{X}_n = \text{Normal}\left(\mu, \frac{1}{n}\right)$$

$$X_{n+1} - \bar{X}_n = \text{Normal}\left(0, 1 + \frac{1}{n}\right)$$

b) 90% of time,

$$-1.645 \leq \underbrace{\frac{X_{n+1} - \bar{X}_n}{\sqrt{1 + \frac{1}{n}}}}_{\substack{|| \\ z}} \leq 1.645$$

$$\text{So } \bar{X}_n - 1.645 \sqrt{1 + \frac{1}{n}} \leq X_{n+1} \leq \bar{X}_n + 1.645 \sqrt{1 + \frac{1}{n}}$$

Here $\bar{X} = 4$; so 90% con int for X_{n+1} is

$$\left(4 - 1.645 \sqrt{1 + \frac{1}{n}}, 4 + 1.645 \sqrt{1 + \frac{1}{n}} \right)$$

7.15) Want $\mu \leq C$ with 99% confidence

So want lower confidence interval

$$\left(-\infty, \bar{X} + \cancel{2.539} t_{19, .01} \frac{S}{\sqrt{20}} \right)$$

$$\begin{aligned} \text{So } C &= \bar{X} + t_{19, .01} \frac{S}{\sqrt{20}} \\ &= 1.2 + 2.539 \frac{\sqrt{.04}}{\sqrt{20}} \end{aligned}$$

$$7.14) \quad \bar{X} \pm t_{19, .005} \frac{S}{\sqrt{n}} \quad \text{or}$$

$$1.2 \pm 2.861 \frac{\sqrt{.04}}{\sqrt{20}}$$

$$7.17) \quad \left. \begin{array}{l} \bar{X} = 333.58 \\ S^2 = 55.18 \end{array} \right\} \text{calculated with excel}$$

$$95\% \text{ confidence int: } \bar{X} \pm t_{23, .025} \frac{S}{\sqrt{24}}$$

$$99\% \text{ confidence int: } \bar{X} \pm t_{23, .005} \frac{S}{\sqrt{24}}$$

$$7.21) \quad 95\% \text{ con int} : \bar{X} \pm t_{99, .025} \frac{16}{\sqrt{100}} \\ = 320 \pm 1.96 \times 1.6$$

$$7.26) \quad \left. \begin{array}{l} \bar{X} = 2062.75 \\ S^2 = 10887.57 \end{array} \right\} \text{ Calculated from excel}$$

$$a) \quad 95\% \text{ con-int} : \bar{X} \pm t_{19, .025} \frac{S}{\sqrt{20}}$$

$$b) \quad 99\% \text{ con int} : \bar{X} \pm t_{19, .005} \frac{S}{\sqrt{20}}$$

c) Want v so that $v \leq \mu$. So v will be the left endpoint of a one-sided upper confidence interval

$$v = \bar{X} - t_{19, .05} \frac{S}{\sqrt{20}}$$